

# Frontier Topics in Empirical Economics: Week 11

## Standard Error Issues

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# Introduction: Nonstandard Standard Error Issues

- Inference is important in practice: Data  $\Rightarrow$  Target distribution
- How accurate is our estimate? How confident are we on our results?
- In traditional inference, we have two assumptions:
  - Uncertainty comes from random-sampling, asymptotics when  $n \rightarrow \infty$
  - i.i.d. sample, no correlations
- What if these two assumptions are violated?

# Introduction: Nonstandard Standard Error Issues

- In this lecture, we consider two cases
- First, when  $n$  is naturally limited (e.g. number of provinces)
- Another type of uncertainty becomes important: Design-based uncertainty
- Second, when i.i.d. fails and errors are clustered
- We have to incorporate this structure in inference
- Angrist calls them "Nonstandard Standard Error Issues"

# Design-based Uncertainty

- In usual case, when we talk about inference, what is that?
- We have a target parameter: "*estimand*"  $\beta$  (**Target**)
- We want to recover it using an "*estimator*" (**Method**)  $\hat{\beta}$  with a sample from the population, which gives you a result called "*estimate*"  $\hat{\beta} = 0.5$  (**Result**)
- This process is called *estimation*, or statistical inference (**Process**)

# Design-based Uncertainty

- Usually, we consider sampling-based uncertainty
- Each time you draw a new sample, it gives you a new estimate from your estimation process
- When sample changes, your estimation result changes
- Uncertainty comes from sampling process
- Thus, you have a standard error for your estimation
- But is this the only uncertainty in empirical research?
- Today, we are going to introduce the second source of uncertainty

# Design-based Uncertainty

- Design-based uncertainty, introduced by Abadie et al. (2020)
- It is the uncertainty coming from the treatment assignment
- Treatment  $X_i$  is no longer considered fixed
- In some cases, person 1 is treated; in other cases, person 1 is not treated
- The potential outcome you observed is different when treatment is randomly changed
- We show that this helps you to understand uncertainty of estimation when you have non-negligible sample size

# Design-based Uncertainty

- To visually explain the difference between traditional sampling-based uncertainty and design-based uncertainty
- Let's take a look at two tables from Abadie et al. (2020)
- $R_i$  is an indicator of whether this observation is included in the sample

- Sampling-based uncertainty

TABLE I  
SAMPLING-BASED UNCERTAINTY (✓ IS OBSERVED, ? IS MISSING)

Unit	Actual Sample			Alternative Sample I			Alternative Sample II			...
	$Y_i$	$Z_i$	$R_i$	$Y_i$	$Z_i$	$R_i$	$Y_i$	$Z_i$	$R_i$	...
1	✓	✓	1	?	?	0	?	?	0	...
2	?	?	0	?	?	0	?	?	0	...
3	?	?	0	✓	✓	1	✓	✓	1	...
4	?	?	0	✓	✓	1	?	?	0	...
⋮	⋮	⋮	⋮	⋮	⋮	⋮	⋮	⋮	⋮	...
$n$	✓	✓	1	?	?	0	?	?	0	...



# Design-based Uncertainty

- Design-based uncertainty

TABLE II  
DESIGN-BASED UNCERTAINTY (✓ IS OBSERVED, ? IS MISSING)

Unit	Actual Sample			Alternative Sample I			Alternative Sample II			...
	$Y_i^*(1)$	$Y_i^*(0)$	$X_i$	$Y_i^*(1)$	$Y_i^*(0)$	$X_i$	$Y_i^*(1)$	$Y_i^*(0)$	$X_i$	...
1	✓	?	1	✓	?	1	?	✓	0	...
2	?	✓	0	?	✓	0	?	✓	0	...
3	?	✓	0	✓	?	1	✓	?	1	...
4	?	✓	0	?	✓	0	✓	?	1	...
⋮	⋮	⋮	⋮	⋮	⋮	⋮	⋮	⋮	⋮	...
$n$	✓	?	1	?	✓	0	?	✓	0	...

# Design-based Uncertainty

- Sampling-based uncertainty
  - Treatment is fixed, sampling observation is random
  - For non-sampled individuals, we cannot observe anything
  - Source of uncertainty: in each sample, we have different observations
- Design-based uncertainty
  - Treatment is random, sampling observation is fixed (e.g. all provinces in China)
  - For each individual, we only observe potential outcome in the realized status (but not counterfactual status)
  - Source of uncertainty: in each sample, we have different treatment status for each individual

# Design-based Uncertainty

- Next, the authors construct a simple model and make the following four points:
  - 1. Show how design-based uncertainty affects the variance of the regression estimator
  - 2. Show White estimator remains conservative when we consider design-based uncertainty
  - 3. We can derive a finite-population correction for White estimator
  - 4. Discuss two sources of uncertainty and external/internal validity

# Design-based Uncertainty

- Assume that we have a **finite** population of size  $n$
- We randomly sample  $N$  from  $n$
- $R_i \in \{0, 1\}$  as an indicator of whether  $i$  is sampled or not
- There is a random binary treatment regressor  $X_i$
- $n_1, N_1$  are treated,  $n_0, N_0$  are not treated
- We have observed and potential outcome as:

$$Y_i = Y_i^*(X_i) = \begin{cases} Y_i^*(1) & \text{if } X_i = 1, \\ Y_i^*(0) & \text{if } X_i = 0 \end{cases}$$

- Potential outcomes are assumed to be non-stochastic

# Design-based Uncertainty

- We use bold letters to represent vector of the whole sample ( $\mathbf{Y}, \mathbf{Y}_i^*(1), \mathbf{Y}_i^*(0), \mathbf{R}$ )
- We define three estimands as our proposed targets
  - Descriptive estimand: free of  $\mathbf{R}$  and potential outcome (population mean difference)
$$\theta^{descr} = \frac{1}{n_1} \sum_{i=1}^n X_i Y_i - \frac{1}{n_0} \sum_{i=1}^n (1 - X_i) Y_i$$
  - Causal estimand: parameter depending on potential outcome  $\mathbf{Y}_i^*(1), \mathbf{Y}_i^*(0)$ 
$$\theta^{causal, sample} = \frac{1}{N} \sum_{i=1}^n R_i (Y_i^*(1) - Y_i^*(0))$$
$$\theta^{causal} = \frac{1}{n} \sum_{i=1}^n (Y_i^*(1) - Y_i^*(0))$$
- $\theta^{causal, sample}$  is the average causal effect of the current sample
- $\theta^{causal}$  is the average causal effect of the whole population

# Design-based Uncertainty

- When estimating  $\theta^{descr}$ , we do not care about design-based uncertainty  
Nothing about treatment or potential outcome
- When estimating  $\theta^{causal,sample}$ , we do not care about sampling-based uncertainty  
Nothing about sampling process (given current sample)
- When estimating  $\theta^{causal}$ , we do care about both types of uncertainty

# Design-based Uncertainty

- To estimate these estimands, we use a simple OLS regression of  $Y_i$  on  $X_i$  to have:

$$\hat{\theta} = \frac{1}{N_1} \sum_{i=1}^n R_i X_i Y_i - \frac{1}{N_0} \sum_{i=1}^n R_i (1 - X_i) Y_i$$

- Sampling-based uncertainty comes from the randomness of  $\mathbf{R}$
- Design-based uncertainty comes from the randomness of  $\mathbf{X}$
- We further assume that both sampling and treatment assignment are random

- It is shown that:

$$E[\hat{\theta}|\mathbf{X}, N_1, N_0] = \theta^{descr}$$

$$E[\hat{\theta}|\mathbf{R}, N_1, N_0] = \theta^{causal, sample}$$

$$E[\hat{\theta}|N_1, N_0] = \theta^{causal}$$

- Conditioning on treatment,  $\theta$  is unbiased for descriptive estimand
- Conditioning on sampling,  $\theta$  is unbiased for causal sample estimand
- Conditioning on none of them,  $\theta$  is unbiased for causal estimand



# Design-based Uncertainty

- We define the population variances as follows:

$$S_x^2 = \frac{1}{n-1} \sum_{i=1}^n \left( Y_i^*(x) - \frac{1}{n} \sum_{j=1}^n Y_j^*(x) \right)^2, \text{ for } x = 0, 1$$

$$S_\theta^2 = \frac{1}{n-1} \sum_{i=1}^n \left( Y_i^*(1) - Y_i^*(0) - \frac{1}{n} \sum_{j=1}^n (Y_j^*(1) - Y_j^*(0)) \right)^2$$

- $S_x^2$  is the variance of potential outcomes for population
- $S_\theta^2$  is the variance of treatment effect for population

- Based on the defined population variance, we can derive three variances

$$V^{total}(N_1, N_0, n_1, n_0) = \text{var}(\hat{\theta} | N_1, N_0) = \frac{S_1^2}{N_1} + \frac{S_0^2}{N_0} - \frac{S_\theta^2}{n_0 + n_1}$$

$$V^{sampling}(N_1, N_0, n_1, n_0) = E[\text{var}(\hat{\theta} | \mathbf{X}, N_1, N_0) | N_1, N_0] = \frac{S_1^2}{N_1} \left(1 - \frac{N_1}{n_1}\right) + \frac{S_0^2}{N_0} \left(1 - \frac{N_0}{n_0}\right)$$

$$V^{design}(N_1, N_0, n_1, n_0) = E[\text{var}(\hat{\theta} | \mathbf{R}, N_1, N_0) | N_1, N_0] = \frac{S_1^2}{N_1} + \frac{S_0^2}{N_0} - \frac{S_\theta^2}{N_0 + N_1}$$

- Now let's analyze them one by one

# Design-based Uncertainty

- $V^{total}$  is the total variance, considering both sampling-based and design-based uncertainty:  $var(\hat{\theta}|N_1, N_0)$
- It is the **variance we want to capture in inference for causal estimator**
- $V^{sampling}$  is the variance from only sampling-based uncertainty, by conditioning on treatment assignment:  $E[var(\hat{\theta}|\mathbf{X}, N_1, N_0)|N_1, N_0]$
- It is the **variance in inference for descriptive estimator**
- $V^{design}$  is the variance from only design-based uncertainty, by conditioning on current sample:  $E[var(\hat{\theta}|\mathbf{R}, N_1, N_0)|N_1, N_0]$
- It is the **variance in inference for causal sample estimator**

# Design-based Uncertainty

- We have the following expressions of variances

$$V^{total}(N_1, N_0, n_1, n_0) = \text{var}(\hat{\theta} | N_1, N_0) = \frac{S_1^2}{N_1} + \frac{S_0^2}{N_0} - \frac{S_\theta^2}{n_0 + n_1}$$

$$V^{sampling}(N_1, N_0, n_1, n_0) = E[\text{var}(\hat{\theta} | \mathbf{X}, N_1, N_0) | N_1, N_0] = \frac{S_1^2}{N_1} \left(1 - \frac{N_1}{n_1}\right) + \frac{S_0^2}{N_0} \left(1 - \frac{N_0}{n_0}\right)$$

$$V^{design}(N_1, N_0, n_1, n_0) = E[\text{var}(\hat{\theta} | \mathbf{R}, N_1, N_0) | N_1, N_0] = \frac{S_1^2}{N_1} + \frac{S_0^2}{N_0} - \frac{S_\theta^2}{N_0 + N_1}$$

- 1. Generally,  $V^{sampling}$  and  $V^{design}$  cannot be ranked, depending on the sampling rates  $\frac{N}{n}$ . A very large sampling rate means a very small  $V^{sampling}$ .

# Design-based Uncertainty

- We have the following expressions of variances

$$V^{total}(N_1, N_0, n_1, n_0) = \text{var}(\hat{\theta} | N_1, N_0) = \frac{S_1^2}{N_1} + \frac{S_0^2}{N_0} - \frac{S_\theta^2}{n_0 + n_1}$$

$$V^{sampling}(N_1, N_0, n_1, n_0) = E[\text{var}(\hat{\theta} | \mathbf{X}, N_1, N_0) | N_1, N_0] = \frac{S_1^2}{N_1} \left(1 - \frac{N_1}{n_1}\right) + \frac{S_0^2}{N_0} \left(1 - \frac{N_0}{n_0}\right)$$

$$V^{design}(N_1, N_0, n_1, n_0) = E[\text{var}(\hat{\theta} | \mathbf{R}, N_1, N_0) | N_1, N_0] = \frac{S_1^2}{N_1} + \frac{S_0^2}{N_0} - \frac{S_\theta^2}{N_0 + N_1}$$

- 2. When  $n \rightarrow \infty$ ,  $V^{sampling} = V^{total}$

If the population is infinite, then design-based uncertainty is ignorable and traditional inference for causal estimand (without considering design-based uncertainty) is fine

# Design-based Uncertainty

- We have the following expressions of variances

$$V^{total}(N_1, N_0, n_1, n_0) = \text{var}(\hat{\theta} | N_1, N_0) = \frac{S_1^2}{N_1} + \frac{S_0^2}{N_0} - \frac{S_\theta^2}{n_0 + n_1}$$

$$V^{sampling}(N_1, N_0, n_1, n_0) = E[\text{var}(\hat{\theta} | \mathbf{X}, N_1, N_0) | N_1, N_0] = \frac{S_1^2}{N_1} \left(1 - \frac{N_1}{n_1}\right) + \frac{S_0^2}{N_0} \left(1 - \frac{N_0}{n_0}\right)$$

$$V^{design}(N_1, N_0, n_1, n_0) = E[\text{var}(\hat{\theta} | \mathbf{R}, N_1, N_0) | N_1, N_0] = \frac{S_1^2}{N_1} + \frac{S_0^2}{N_0} - \frac{S_\theta^2}{N_0 + N_1}$$

- 3. Consider estimating  $\theta^{descr}$  or  $\theta^{causal}$ :

When population is finite,  $V^{total}$  and  $V^{sampling}$  are overstated if we think it is infinite

$$V^{total}(N_1, N_0, \infty, \infty) - V^{total}(N_1, N_0, n_1, n_0) = \frac{S_\theta^2}{n_0 + n_1} \geq 0,$$

$$V^{sampling}(N_1, N_0, \infty, \infty) - V^{sampling}(N_1, N_0, n_1, n_0) = \frac{S_1^2}{n_1} + \frac{S_0^2}{n_0} \geq 0$$

# Design-based Uncertainty

- We have the following expressions of variances

$$V^{total}(N_1, N_0, n_1, n_0) = \text{var}(\hat{\theta} | N_1, N_0) = \frac{S_1^2}{N_1} + \frac{S_0^2}{N_0} - \frac{S_\theta^2}{n_0 + n_1}$$

$$V^{sampling}(N_1, N_0, n_1, n_0) = E[\text{var}(\hat{\theta} | \mathbf{X}, N_1, N_0) | N_1, N_0] = \frac{S_1^2}{N_1} \left(1 - \frac{N_1}{n_1}\right) + \frac{S_0^2}{N_0} \left(1 - \frac{N_0}{n_0}\right)$$

$$V^{design}(N_1, N_0, n_1, n_0) = E[\text{var}(\hat{\theta} | \mathbf{R}, N_1, N_0) | N_1, N_0] = \frac{S_1^2}{N_1} + \frac{S_0^2}{N_0} - \frac{S_\theta^2}{N_0 + N_1}$$

- 4. Consider estimating  $\theta^{causal, sample}$ :

When population is finite,  $V^{design}$  is fine even if we think it is infinite

$$V^{design}(N_1, N_0, \infty, \infty) = V^{design}(N_1, N_0, n_1, n_0)$$

Relative sample size does not affect variance conditional on current sample

# Design-based Uncertainty

- In practice, we usually use White estimator of the variance matrix
- It is **calculated without considering design-based uncertainty**<sup>1</sup>

$$\hat{V}^w = \frac{\hat{S}_1^2}{N_1} + \frac{\hat{S}_0^2}{N_0}, \text{ where } \hat{S}_1^2 = \frac{1}{N_1 - 1} \sum_{i=1}^n R_i X_i \left( Y_i - \frac{1}{N_1} \sum_{i=1}^n R_i X_i Y_i \right)^2$$

- It is unbiased for  $V^{total}$  when  $n$  is infinite
- The small population bias is  $E[\hat{V}^w | N] - V^{total} = S_\theta^2 / n$

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<sup>1</sup> $\hat{S}_0^2$  is defined analogously



# Design-based Uncertainty

- We can see that if we ignore design-based uncertainty in inference
- It is fine if we have a small sample compared with a massive population
- Like you have a CFPS dataset to represent all families in China
- But the positive bias will become large if we have a large sample size compared with a limited population
- Like you have a province-level regression
- In this case, traditional variance estimation can be too large and too conservative
- Because you ignore the fact that you already have a large part of the population

# Design-based Uncertainty

- But fortunately, we can derive a bias-corrected estimator
- By taking into consideration
  - You have a large sample relative to a small population
  - You have uncertainty in treatment assignment
- The derivation of this estimator is technical
- Read Abadie et al. (2020) if you are interested

# Clustered Standard Errors: Motivating Example

- Next, let's consider the clustering issue
- Many scholars claim that smaller classes are better
- What is the impact of class size on students' achievement?
- Hard to identify using observational data (selection problem)
- STAR is a RCT to answer this question

## Clustered Standard Errors: Motivating Example

- It involves 11,600 children in TN
- Kids are randomly assigned to two kinds of classes
  - (1) Small class with 13-17 children; (2) Regular class with 22-25 children
- Then we can identify the treatment effect of class size
- One assumption we always make is i.i.d.
- However, students in the same class are of course not independently sampled
- What will happen if we have correlations at class/school/district... level?

# Clustered Standard Errors: Motivating Example

- The short answer is: we may underestimate the standard error
- Let's see why it is and how to fix this issue

## Clustered Standard Errors: Setting

- Let's go on with the STAR experiment
- Consider the following regression for student  $i$  in class  $g$ :

$$y_{ig} = \beta_0 + \beta_1 x_g + e_{ig}$$

- $y_{ig}$  test score;  $x_g$  class size (randomly assigned);  $e_{ig}$  error term
- This is a special case when  $x$  is fixed at  $g$  level (same treatment for the whole class)
- Test scores in the same class tend to be correlated (Same environment, teacher...)

## Clustered Standard Errors: Setting

- Thus, we give up i.i.d. assumption and assume that for student  $i$  and  $j$ :

$$E[e_{ig}e_{jg}] = \rho_e \sigma_e^2 > 0$$

- $\rho_e$  is the error intraclass correlation,  $\sigma_e^2$  is the error variance
- Assume that we can decompose error into

$$e_{ig} = \nu_g + \eta_{ig}, \quad \nu_g \perp\!\!\!\perp \eta_{ig}$$

- We assume that  $\nu_g$  captures all within class correlations ( $\eta_{ig} \perp\!\!\!\perp \eta_{jg}$ )
- Also assume homoskedasticity for both  $\nu_g$  and  $\eta_{ig}$
- Then we can prove that

$$\rho_e = \frac{\sigma_\nu^2}{\sigma_\nu^2 + \sigma_\eta^2} \tag{1}$$

- Intraclass correlation is the share of intraclass uncertainty in the total uncertainty

## Clustered Standard Errors: Setting

- Equation (1) is called "intracluster correlation coefficient"
- Homework: Derive equation (1) from the previous setting



## Clustered Standard Errors: Bias and Moulton Factor

- Let  $V_c(\hat{\beta}_1)$  be the conventional OLS variance,  $V(\hat{\beta}_1)$  be the correct variance
- Assume we have classes with equal size  $n$ , then

$$\frac{V(\hat{\beta}_1)}{V_c(\hat{\beta}_1)} = 1 + (n - 1)\rho_e$$

- We call this Moulton factor
- $n, \rho_e \uparrow \Rightarrow$  Bias of conventional variance  $\uparrow$
- Larger  $n$  means fewer groups  $\Rightarrow$  less information
- Homework 2: What will happen if  $\rho_e = 1$ ? (Answer in MHE)

## Clustered Standard Errors: Bias and Moulton Factor

- Previous setting assumes fixed  $x_g$  within each group
- Let's see Moulton factor in a more general case when  $x_{ig}$  can vary across  $i$  in the same group

$$\frac{V(\hat{\beta}_1)}{V_c(\hat{\beta}_1)} = 1 + \left[ \frac{V(n_g)}{\bar{n}} + \bar{n} - 1 \right] \rho_x \rho_e \quad (2)$$

- $\bar{n}$  is average group size;  $V(n_g)$  is variance of group sizes;  $\rho_x$  is intraclass correlation of  $x_{ig}$

# Clustered Standard Errors: Bias and Moulton Factor

- In general, bias from within class correlation is larger when
  - (1) Average group size  $\uparrow$
  - (2) Variance of group size  $\uparrow$
  - (3) Intraclass correlation of treatment  $x_{ig}$   $\uparrow$
  - (4) Error intraclass correlation  $\uparrow$
- The implication of (3)
  - Bias can be very large in the fixed group treatment  $x_g$  case
  - No need to cluster anything if the assignment is totally random for every individual
- The implication of (4): Naturally, no bias when  $\rho_e = 0$

## Clustered Standard Errors: Fix the Bias

- Now we know that std error estimation can be biased when we have correlation within classes
- What we should do? Several methods are available
  - (1) Use Moulton factor equation (2) to correct  
Not that good: error structure assumptions (homoskedasticity)
  - (2) **Recommended: Liang and Zeger (1986) clustering estimator**  
Generally consistent as number of groups  $\rightarrow \infty$  (In stata, use option *cluster*)
  - (3) Running group-level regressions  $\bar{y}_g = \beta_0 + \beta_1 x_g + \bar{e}_g$  using WLS (group size as weights)  
Better finite-sample properties, but  $x_g$  has to be group-fixed
  - Other methods: Block bootstrap, MLE...

## Clustered Standard Errors: Choosing Cluster Levels

- How to choose the level of clustering?
- In STAR experiment, why not boy/girl, black/white/asian...?
- Clustering in more dimensions/higher level gives you larger std errs
- Is that OK to always cluster in more and more dimensions (be conservative)? NO.  
You can be too conservative  $\Rightarrow$  Overestimate std err
- Similarly, not always good to cluster at higher and higher level

## Clustered Standard Errors: Choosing Cluster Levels

- This is because when you cluster in more and more dimensions
- Or at higher and higher level
- Your effective sample size compared with effective population becomes larger and larger
- As Abadie et al. (2020) has shown, it leads to overestimation of the std err
- For example, you have data of 10,000 firms in 20 provinces
- 10,000 can be a very small proportion of all firms in mainland China
- When you cluster at province level, effective sample rate becomes  $20/31!$

# Clustered Standard Errors: Choosing Cluster Levels

- Thus, two issues remains
  - How to choose cluster level reasonably?
  - How to incorporate design-based uncertainty?
- Abadie et al. (2023) considers clustering as a sampling/design problem
- Cluster level depends on how you get your samples/assign your treatment
- It comes from the basic idea of Abadie et al. (2020)
- You have to consider both sampling-based and design-based uncertainty
- This is more to the core of the clustering problem

# Clustered Standard Errors: Choosing Cluster Levels

- There are three misconceptions they want to clarify
- 1. The need for clustering hinges on the presence of a correlation between residuals
  - No. The essence is the clustering of sampling or treatment assignment
  - Even if students' scores are correlated within classroom, there is no need to cluster when sampling and treatment are totally random
- 2. No harm in using clustered std err when they are not required
  - Confidence intervals will be unnecessarily conservative
- 3. Researchers either fully adjust for clustering by using Liang and Zeger (1986) or not do that at all
  - Not really. They propose a new estimator CCV/TSCB to correct for large effective sample rate in clustering



# Clustered Standard Errors: Choosing Cluster Levels

Here are some empirical suggestions from Abadie et al. (2023)

- 1. If sampling and treatment are both random
  - Do not cluster!
  - In this case, if sample represents a large fraction of the population, even White estimator is too conservative (Abadie et al., 2020)
- 2. If random sampling but clustered treatment assignment
  - Cluster at the treatment level.
  - In the fuzzy design case, using CCV/TSCB estimator

# Clustered Standard Errors: Choosing Cluster Levels

- 3. If clustered sampling, random treatment assignment
  - Cluster at the sampling level, if you have small fraction of sampled clusters or small fraction of sampled units within each cluster
  - This is specifically important in panel data analysis
  - Do not cluster in other cases
- 4. If clustered sampling, clustered treatment assignment
  - Cluster at the higher level to be conservative

# Clustered Standard Errors: Choosing Cluster Levels

- Let us go over two practical examples
- Case 1: (*Sampling cluster*) Some household/firm survey will
  - (1) Randomly select 50/300 cities in China
  - (2) Randomly select 100 households in each sampled city
- It gives you a natural stratified data set
- Just cluster at city level (in general, first sampling stage level)
- Case 2: (*Treatment cluster*) STAR assigns treatment at class level
- Then just cluster at class level

# Clustered Standard Errors: DID and Serial Correlation

- One special case we must underscore is panel data analysis
- When using panel data, we usually employ time variation for identification
- You draw people, but not people in a specific year  $\Rightarrow$  serial correlation
- You are drawing samples/assign treatment clustered at individual level
- Thus, DID gives a natural clustering structure of error
- **One-level-up principle:**  
Cluster at individual/province/city level, but NEVER individual-year/province-year/city-year level!!

# Conclusion

- Today we discuss two nonstandard standard error issues
  - When sample is large compared with population
  - When errors are not i.i.d. but clustered
- In the first issue, we claim that we need to consider both sampling-based and design-based uncertainty
- Using traditional inference will have too large and conservative std err

# Conclusion

- In the second case, we find that not adjusting for cluster will generate a too small std err
- We can use LZ estimator to fix it (consistent as  $\#groups \rightarrow \infty$ )
- Clustering at higher level is not always good
- Clustering comes from either clustered sampling or clustered treatment
- Cluster at the first sampling stage, or treatment assignment level
- Do NOT cluster if you have a totally random sample and random treatment
- In DID, cluster one level up to take care of the serial correlation

# References

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